10: Normal Distributions

Jerry Cain April 22nd, 2024

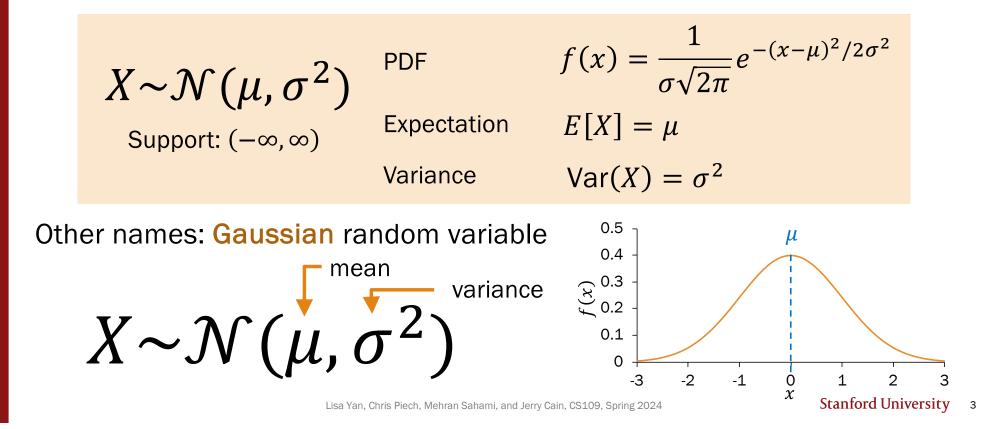
Lecture Discussion on Ed

Normal Random Variables



Normal Random Variable

<u>def</u> A Normal random variable *X* is defined as follows:



Carl Friedrich Gauss

Carl Friedrich Gauss (1777-1855) was a remarkably influential German mathematician.



Johann Carl Friedrich Gauss (/gaus/; German: Gauß [gaus] (•) listen); Latin: Carolus Fridericus Gauss; 30 April 1777 – 23 February 1855) was a German mathematician and physicist who made significant contributions to many fields, including algebra, analysis, astronomy, differential geometry, electrostatics, geodesy, geophysics, magnetic fields, matrix theory, mechanics, number theory, optics and statistics. Sometimes referred to as the *Princeps mathematicorum*^[1] (Latin for "the foremost of mathematicians") and "the greatest mathematician since antiquity", Gauss had an exceptional influence in many fields of mathematics and science, and is ranked among history's most influential mathematicians.^[2]

Did **not** invent Normal distribution but rather popularized it.

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Why the Normal?

- Common for natural phenomena: height, weight, etc.
- Most noise in the world is Normal
- Often results from the sum of many random variables
- Sample means are distributed normally

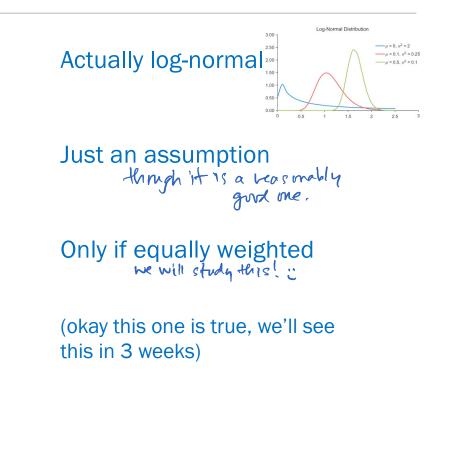
That's what they want you to believe...



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Why the Normal?

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Okay, so why the Normal?

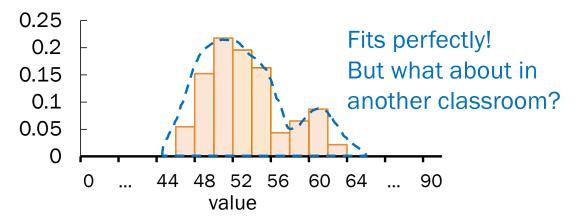
Part of CS109 learning goals:

• Translate a problem statement into a random variable

In other words: model real life situations with probability distributions

How do you model student heights?

• Suppose you have data from one classroom.



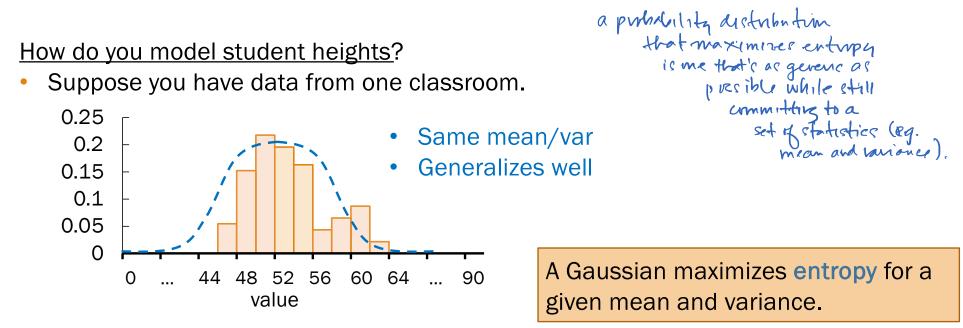
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Okay, so why the Normal?

Part of CS109 learning goals:

• Translate a problem statement into a random variable

In other words: model real life situations with probability distributions



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Why the Normal?

- because it's well understood Common for natural phenomena: height, weight, etc.
- Most noise in the world is Normal
- Often results from random va

Samp

(okay this one is true, we'll see this in 3 weeks)

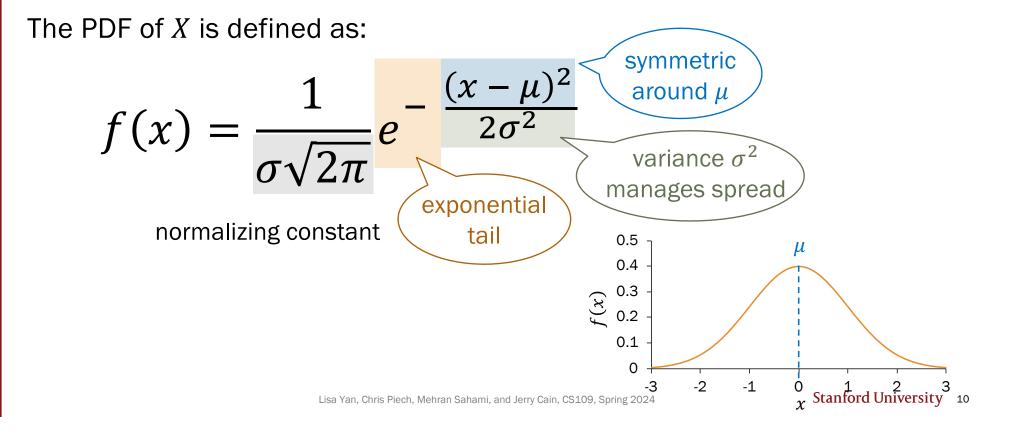
Actually log-norr

Stay critical of how to model realworld phenomena.

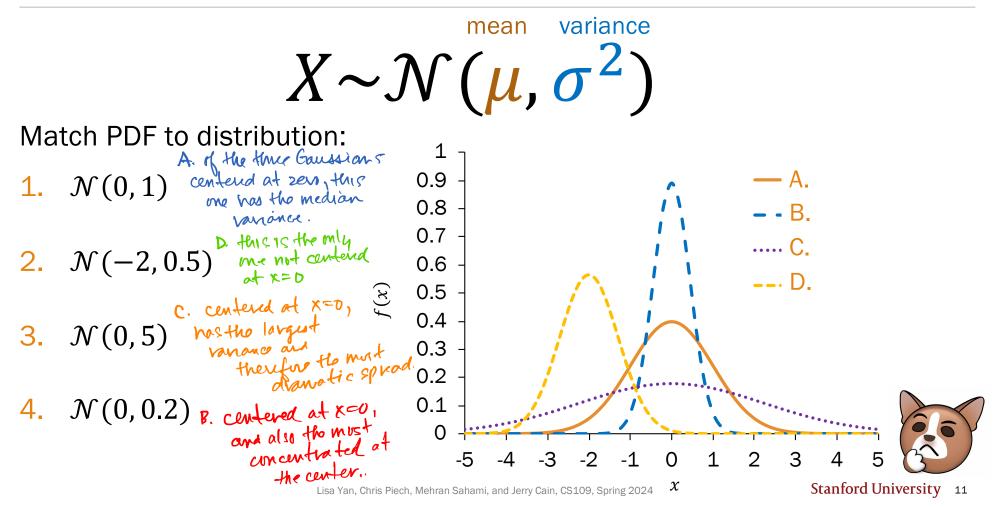
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Anatomy of a beautiful equation

Let $X \sim \mathcal{N}(\mu, \sigma^2)$.



Normal Random Variable

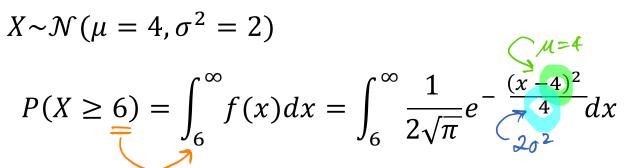


Getting to class

You spend some minutes, *X*, traveling between classes.

- Average time spent: $\mu = 4$ minutes
- Variance of time spent: $\sigma^2 = 2 \text{ minutes}^2$

Suppose X is normally distributed. What is the probability you spend ≥ 6 minutes traveling?



(tell Jerry if you solve this analytically and we'll be famous together)

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Love and Anger in the) Same Formula

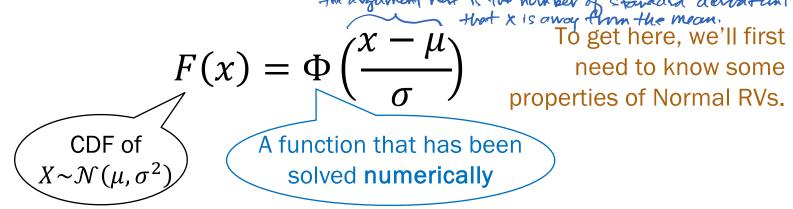


Computing probabilities with Normal RVs

For a Normal RV $X \sim \mathcal{N}(\mu, \sigma^2)$, its CDF has no closed form.

$$P(X \le x) = F(x) = \int_{-\infty}^{x} \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(y-\mu)^2}{2\sigma^2}} dy$$
 Cannot be solved analytically

However, we can solve for probabilities numerically using a function Φ :



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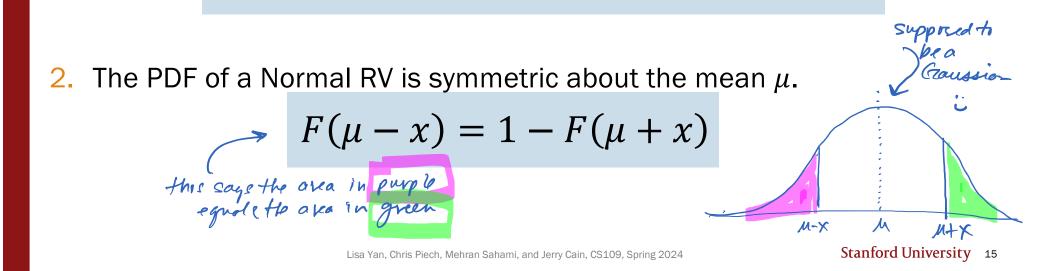
Normal RV: Properties

Properties of Normal RVs

Let $X \sim \mathcal{N}(\mu, \sigma^2)$ with CDF $P(X \leq x) = F(x)$.

1. Linear transformations of Normal RVs are also Normal RVs.

If
$$Y = aX + b$$
, then $Y \sim \mathcal{N}(a\mu + b, a^2\sigma^2)$.



1. Linear transformations of Normal RVs

Let
$$X \sim \mathcal{N}(\mu, \sigma^2)$$
 with CDF $P(X \le x) = F(x)$.

Linear transformations of X are also Normal.

If
$$Y = aX + b$$
, then $Y \sim \mathcal{N}(a\mu + b, a^2\sigma^2)$

Proof:

•
$$E[Y] = E[aX + b] = aE[X] + b = a\mu + b$$
 Linearity of Expectation

•
$$\operatorname{Var}(Y) = \operatorname{Var}(aX + b) = a^2 \operatorname{Var}(X) = a^2 \sigma^2 \operatorname{Var}(aX + b) = a^2 \operatorname{Var}(X)$$

• Y is also Normal

Proof in Ross, 10th ed (Section 5.4)

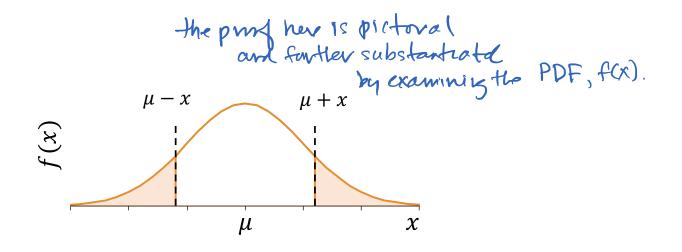
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2. Symmetry of Normal RVs

Let
$$X \sim \mathcal{N}(\mu, \sigma^2)$$
 with CDF $P(X \le x) = F(x)$.

The PDF of a Normal RV is symmetric about the mean μ .

$$F(\mu - x) = 1 - F(\mu + x)$$



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Using symmetry of the Normal RV

Let $Z \sim \mathcal{N}(0,1)$ with CDF $P(Z \leq z) = F(z)$.

Suppose we only knew numeric values for F(z) and F(y), for some $y, z \ge 0$.

How do we compute the following probabilities?

F(z)

1.
$$P(Z \le z)$$
 =

 2. $P(Z < z)$
 =

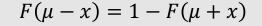
 3. $P(Z \ge z)$
 =

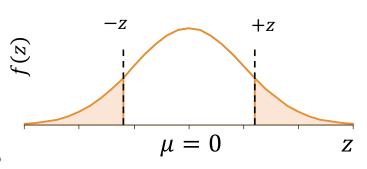
 4. $P(Z \le -z)$
 =

 5. $P(Z \ge -z)$
 =

 6. $P(y < Z < z)$

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A.
$$F(z)$$

B. $1 - F(z)$
C. $F(z) - F(y)$



Using symmetry of the Normal RV

Let $Z \sim \mathcal{N}(0,1)$ with CDF $P(Z \leq z) = F(z)$.

Suppose we only knew numeric values for F(z) and F(y), for some $y, z \ge 0$.

How do we compute the following probabilities?

1. $P(Z \leq z)$ **2.** P(Z < z)3. $P(Z \ge z)$ $4. \quad P(Z \leq -z)$ 5. $P(Z \ge -z) = F(z)$ 6. P(y < Z < z)

$$= F(z)$$

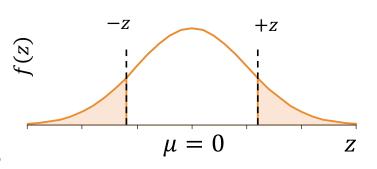
 $= F(z)$

$$= F(z)$$
$$= 1 - F(z)$$

= 1 - F(z)

- = F(z) F(y)

 $F(\mu - x) = 1 - F(\mu + x)$



| Α. | F(z) |
|----|-------------|
| Β. | 1-F(z) |
| C. | F(z) - F(y) |

Symmetry is particularly useful when computing probabilities of zero-mean Normal RVs.

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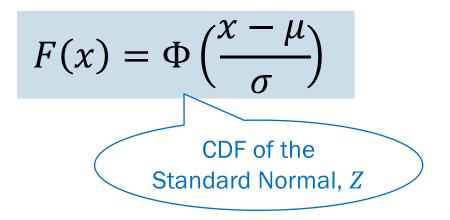
Normal RV: Computing probability

Computing probabilities with Normal RVs

Let $X \sim \mathcal{N}(\mu, \sigma^2)$.

To compute the CDF, $P(X \le x) = F(x)$:

- We cannot analytically solve the integral, as it has no closed form.
- ... but we *can* solve numerically using a function Φ :



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Standard Normal RV, Z

The **Standard Normal** random variable *Z* is defined as follows:

 $Z \sim \mathcal{N}(0, 1)$

Expectation E[Z]Variance Var(

$$E[Z] = \mu = 0$$

Var(Z) = $\sigma^2 = 2$

Note: not a new distribution; just a special case of the Normal

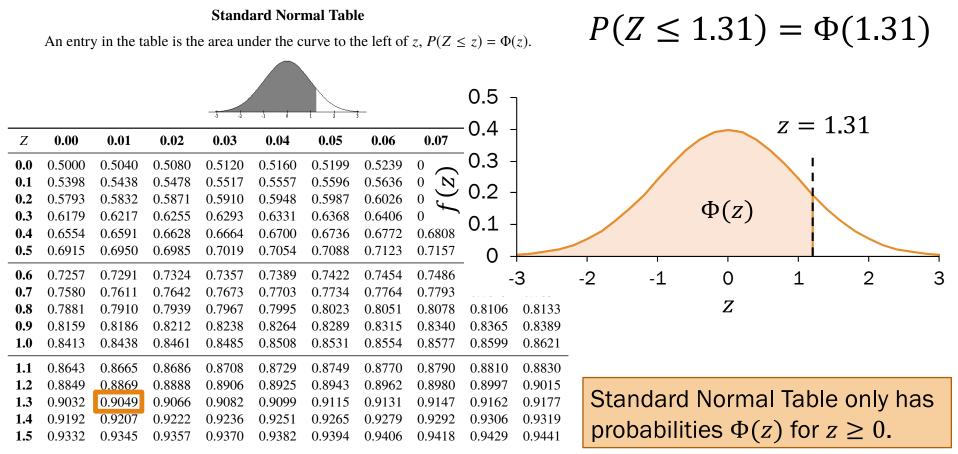
Λ

Other names: Unit Normal

CDF of Z defined as: $P(Z \le z) = \Phi(Z)$

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Φ has been numerically computed



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History fact: Standard Normal Table

TABLES

SERVANT

AU CALCUL DES RÉFRACTIONS APPROCHANTES DE L'HORIZON.

TABLE PREMIÈRE.

Intégrales de e^{-it} dt, depuis une valeur quelconque de t jusqu'à t infinie,

| 1 | Intégrale. | Diff. prem. | Diff. II. | Diff. III. |
|------|------------|-------------|-----------|------------|
| 0,00 | 0,88622692 | 999968 | 201 | 199 |
| 0,01 | 0,87622724 | 999767 | 400 | 199 |
| 0.02 | 0.86622057 | 999367 | 599 | 200 |
| 0,03 | 0,85623590 | 998768 | 799 | 199 |
| 0,04 | 0,84624822 | 997969 | 998 | 197 |
| 0,05 | 0,83626853 | 99697 I | 1195 | 199 |
| 0,06 | 0,82629882 | 995776 | 1394 | 196 |

The first Standard Normal Table was computed by Christian Kramp, French astronomer (1760–1826), in Analyse des Réfractions Astronomiques et Terrestres, 1799

Used a Taylor series expansion to the third power

| integral from x = 0.03 to infinity of e^{-x^2} | | | | |
|--|----------|--|--|--|
| $\int_{\Sigma^{2}}^{\pi}$ Extended Keyboard | 1 Upload | | | |
| Definite integral: | | | | |
| $\int_{0.03}^{\infty} e^{-x^2} dx = 0.856236$ | | | | |

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Probabilities for a general Normal RV

Let $X \sim \mathcal{N}(\mu, \sigma^2)$. To compute the CDF $P(X \leq x) = F(x)$, we use Φ , the CDF for the Standard Normal $Z \sim \mathcal{N}(0, 1)$: $F(x) = \Phi\left(\frac{x-\mu}{x-\mu}\right)$ $= P(X \le x)$ $= P(X - \mu \le x - \mu) = P\left(\frac{X - \mu}{\sigma} \le \frac{x - \mu}{\sigma}\right)$ $= P(X - \mu \le x - \mu) = P\left(\frac{X - \mu}{\sigma} \le \frac{x - \mu}{\sigma}\right)$ Algebra + $\sigma > 0$ Proof: $F(x) = P(X \le x)$ $= P\left(Z \le \frac{x-\mu}{\sigma}\right) \quad \left\{\begin{array}{l} \cdot \frac{X-\mu}{\sigma} = \frac{1}{\sigma}X - \frac{\mu}{\sigma} \text{ is a linear transform of } X.\\ \cdot \text{ This is distributed as } \mathcal{N}\left(\frac{1}{\sigma}\mu - \frac{\mu}{\sigma}, \frac{1}{\sigma^2}\sigma^2\right) = \mathcal{N}(0,1)\\ \cdot \text{ In other words, } \frac{X-\mu}{\sigma} = Z \sim \mathcal{N}(0,1) \text{ with CDF } \Phi.\end{array}\right.$

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Probabilities for a general Normal RV

Let $X \sim \mathcal{N}(\mu, \sigma^2)$. To compute the CDF $P(X \le x) = F(x)$, we use Φ , the CDF for the Standard Normal $Z \sim \mathcal{N}(0, 1)$:

$$F(x) = \Phi\left(\frac{x-\mu}{\sigma}\right)$$

Proof:

$$F(x) = P(X \le x)$$

$$= P(X - \mu \le x - \mu) = P\left(\frac{X - \mu}{\sigma} \le \frac{x - \mu}{\sigma}\right)$$
Definition of CDF
Algebra + $\sigma > 0$

$$= P\left(Z \le \frac{x - \mu}{\sigma}\right) - \left[\underbrace{\cdot \frac{X - \mu}{\sigma} = \frac{1}{\sigma} X - \frac{\mu}{\sigma}}_{\sigma} \text{ is a linear transform of } X. \underbrace{\cdot \frac{1}{\sigma} = \frac{1}{\sigma} X - \frac{\mu}{\sigma}}_{\sigma} \text{ is a linear transform of } X.$$

$$= \Phi\left(\frac{x - \mu}{\sigma}\right) - \left[\underbrace{\cdot \frac{X - \mu}{\sigma} = \frac{1}{\sigma} X - \frac{\mu}{\sigma}}_{c} \text{ is a linear transform of } X. \underbrace{\cdot \frac{1}{\sigma} = \frac{1}{\sigma} X - \frac{\mu}{\sigma}}_{c} \text{ is a linear transform of } X.$$

$$= \Phi\left(\frac{x - \mu}{\sigma}\right) - \left[\underbrace{\cdot \frac{1}{\sigma} = \frac{1}{\sigma} X - \frac{\mu}{\sigma}}_{c} \text{ is a linear transform of } X. \underbrace{\cdot \frac{1}{\sigma} = \frac{1}{\sigma} X - \frac{\mu}{\sigma}}_{c} \text{ is a linear transform of } X.$$

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Campus bikes

You spend some minutes, X, traveling between classes.

- Average time spent: $\mu = 4$ minutes Variance of time spent: $\sigma^2 = 2$ minutes² •

Suppose X is normally distributed. What is the probability you spend \geq 6 minutes traveling?

table

$$X \sim \mathcal{N}(\mu = 4, \sigma^2 = 2)$$
 $X \sim P(X \ge 6) = \int_6^\infty f(x) dx$ (no analytic solution)

- 00

1. Compute
$$z = \frac{(x-\mu)}{\sigma}$$

 $P(X \ge 6) = 1 - F_x(6)$
 $= 1 - \Phi\left(\frac{6-4}{\sqrt{2}}\right)$
 $\approx 1 - \Phi(1.41)$
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Is there an easier way? (yes)

Let $X \sim \mathcal{N}(\mu, \sigma^2)$. What is $P(X \le x) = F(x)$?

Use Python

```
from scipy import stats
X = stats.norm(mu, std)
X.cdf(x)
```

SciPy reference: https://docs.scipy.org/doc/scipy/refere nce/generated/scipy.stats.norm.html

I'm not sure why Python decided to parameterize **stats.norm** around the standard deviation instead of the variance, but it did. ^(C)

Exercises

Let $X \sim \mathcal{N}(\mu = 3, \sigma^2 = 16)$. Std deviation $\sigma = 4$. If $X \sim \mathcal{N}(\mu, \sigma^2)$, then $F(x) = \Phi\left(\frac{x-\mu}{\sigma}\right)$ 1. P(X > 0) = I - P(X < D)= 1 - F(0)Symmetry of the PDF of Normal RV implies $= 1 - \phi\left(\frac{0-3}{4}\right)$ $\Phi(-z) = 1 - \Phi(z)$ $= 1 - \phi(\frac{-3}{4})$ $= (l - \phi(\frac{3}{4}))$ $= \phi\left(\frac{3}{4}\right) = 0.7734$ best to use Pythm's scipy package to Ink this up . :

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Let $X \sim \mathcal{N}(\mu = 3, \sigma^2 = 16)$. Note standard deviation $\sigma = 4$.

How would you write each of the below probabilities as a function of the standard normal CDF, Φ ?

P(X > 0)
 P(2 < X < 5)
 P(|X - 3| > 6)

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- If $X \sim \mathcal{N}(\mu, \sigma^2)$, then $F(x) = \Phi\left(\frac{x-\mu}{\sigma}\right)$
- Symmetry of the PDF of Normal RV implies $\Phi(-z) = 1 - \Phi(z)$



Let $X \sim \mathcal{N}(\mu = 3, \sigma^2 = 16)$. Std deviation $\sigma = 4$. 1. P(X > 0)2. P(2 < X < 5) = F(5) - F(2) $= \Phi(\frac{5-3}{4}) - \Phi(\frac{2-3}{4})$ $= \Phi(\frac{1}{2}) - \Phi(\frac{1}{4})$ $= \Phi(\frac{1}{2}) - (1 - \Phi(\frac{1}{4}))$ $= \Phi(\frac{1}{2}) + \Phi(\frac{1}{4}) - 1$ $= \Phi(\frac{1}{2}) + \Phi(\frac{1}{4}) - 1$

• If
$$X \sim \mathcal{N}(\mu, \sigma^2)$$
, then
 $F(x) = \Phi\left(\frac{x-\mu}{\sigma}\right)$

Symmetry of the PDF of Normal RV implies $\Phi(-z) = 1 - \Phi(z)$

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Let $X \sim \mathcal{N}(\mu = 3, \sigma^2 = 16)$. Std deviation $\sigma = 4$. 1. P(X > 0)2. P(2 < X < 5)3. P(|X - 3| > 6)

Compute $z = \frac{(x-\mu)}{\sigma}$ P(X < -3) + P(X > 9) = F(-3) + (1 - F(9)) $= \Phi\left(\frac{-3-3}{4}\right) + \left(1 - \Phi\left(\frac{9-3}{4}\right)\right)$ • If $X \sim \mathcal{N}(\mu, \sigma^2)$, then $F(x) = \Phi\left(\frac{x-\mu}{\sigma}\right)$

• Symmetry of the PDF of Normal RV implies $\Phi(-x) = 1 - \Phi(x)$

Look up $\Phi(z)$ in table

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Let $X \sim \mathcal{N}(\mu = 3, \sigma^2 = 16)$. Std deviation $\sigma = 4$. 1. P(X > 0)2. P(2 < X < 5)3. P(|X - 3| > 6)

Compute
$$z = \frac{(x-\mu)}{\sigma}$$

 $P(X < -3) + P(X > 9)$
 $= F(-3) + (1 - F(9))$
 $= \Phi\left(\frac{-3-3}{4}\right) + \left(1 - \Phi\left(\frac{9-3}{4}\right)\right)$
Look up $\Phi(z)$ in table
 $= \Phi\left(-\frac{3}{2}\right) + \left(1 - \Phi\left(\frac{3}{2}\right)\right)$
 $= 2\left(1 - \Phi\left(\frac{3}{2}\right)\right)$
 ≈ 0.1337

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If $X \sim \mathcal{N}(\mu, \sigma^2)$, then

Symmetry of the PDF of

 $F(x) = \Phi\left(\frac{x-\mu}{\sigma}\right)$

Normal RV implies

 $\Phi(-x) = 1 - \Phi(x)$

•

Noisy Wires

Send a voltage of 2 V or -2 V on wire (to denote 1 and 0, respectively).

- X =voltage sent (2 or -2)
- $Y = \text{noise}, Y \sim \mathcal{N}(0, 1)$
- R = X + Y voltage received.

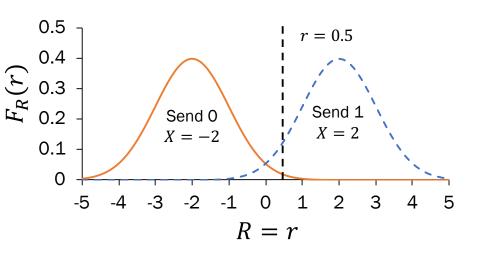
Decode:

1 if $R \ge 0.5$ 0 otherwise.

- What is P(decoding error | original bit is 1)?
 i.e., we sent 1, but we decoded as 0?
- 2. What is P(decoding error | original bit is 0)?

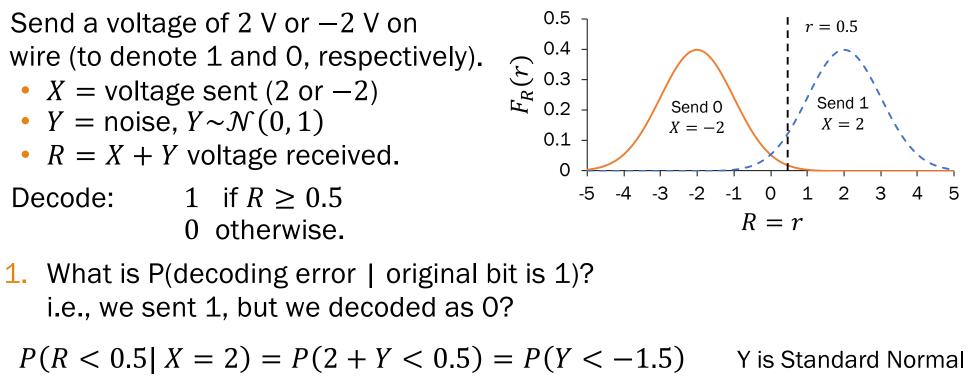
These probabilities are unequal. Why might this be useful?

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Noisy Wires

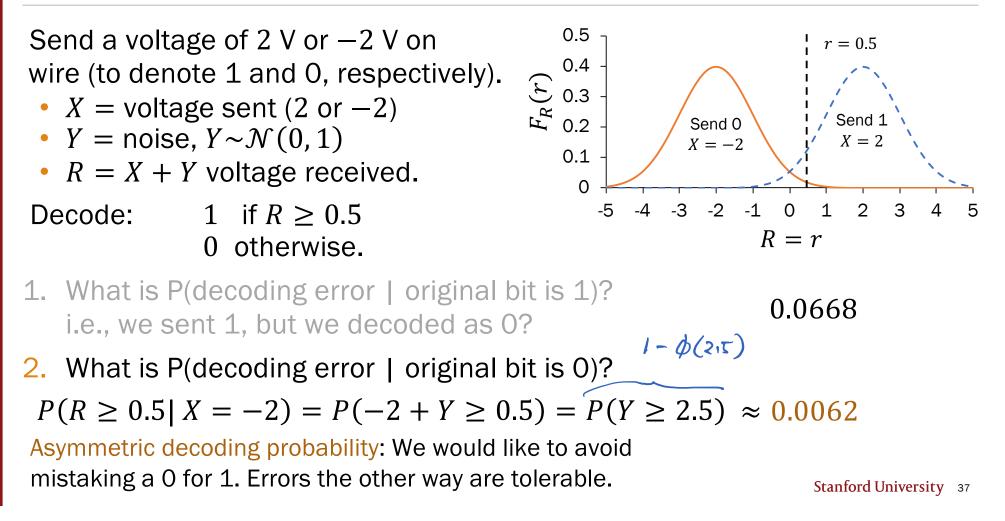


$$= \Phi(-1.5) = 1 - \Phi(1.5) \approx 0.0668$$

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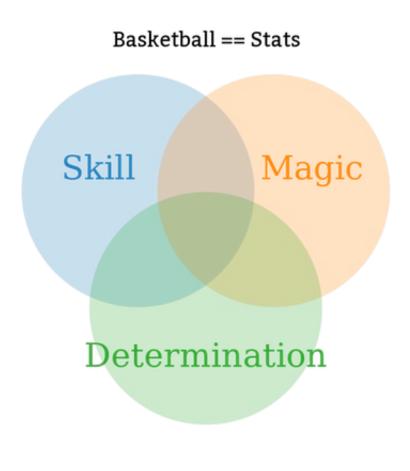
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Noisy Wires



Sampling with the Normal RV

ELO ratings





What is the probability that the Warriors win? More generally: How can you model zero-sum games?

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ELO ratings

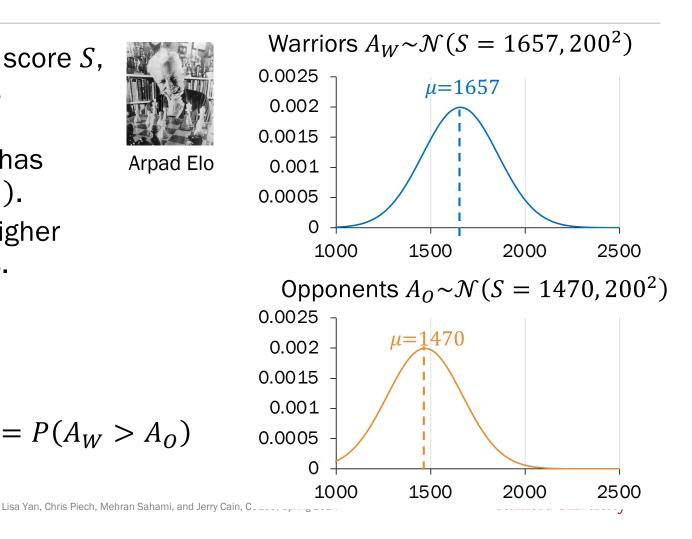
Each team has an ELO score S, calculated based on its past performance.

- Each game, a team has ability $A \sim \mathcal{N}(S, 200^2)$.
- The team with the higher sampled ability wins.

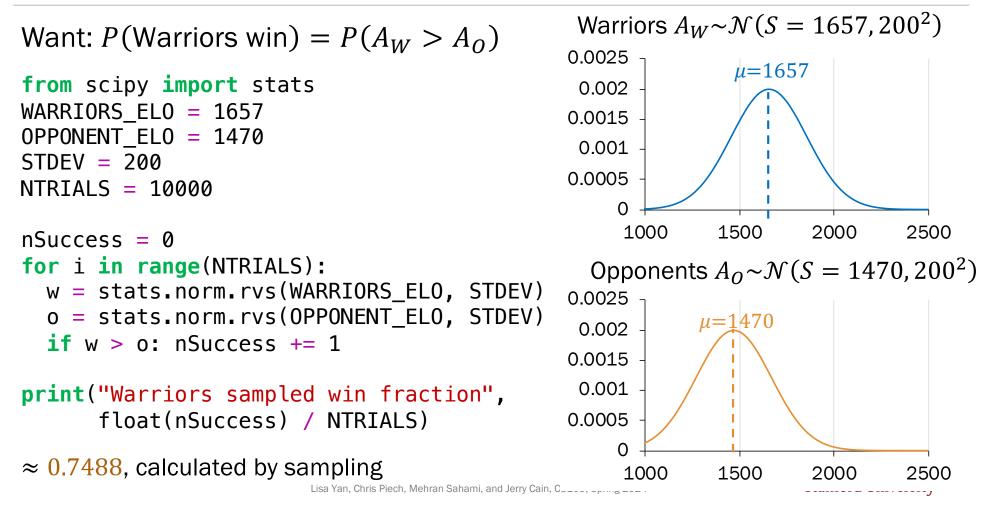
What is the probability that Warriors win this game?

Want: $P(\text{Warriors win}) = P(A_W > A_O)$

Arpad Elo



ELO ratings



Is there a better way?

 $P(A_W > A_O)$



 This is a probability of an event involving two continuous random variables!

actual depiction of someone understanding joint continuous random variables

We'll solve this problem analytically in less than two weeks' time.

Big goal for next lecture: Events involving two discrete random variables. Stay tuned!